

**Multiterminal Estimation -
Extensions and a Geometric Interpretation.**

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Outline

1. Introduction
2. Zhang and Berger
3. Han and Amari
4. Extending Zhang and Berger
5. Geometric interpretation
6. Conclusion

Multiterminal Data Compression

Deals with the problem of optimal inference based on multi-source lossy data.

Multiterminal framework: Provides a measure of **loss of statistical information due to the separate compression of multiple sources.**

- Best case scenario
- Baseline for comparison of methods

Multiterminal Estimation Problem: **Level of compression/loss of statistical estimation efficiency** we **want/can accept?**

Multiterminal Data Compression

- Remote sensing, Signal Detection, Pattern Recognition.

Multiterminal hypothesis testing

- CEO problem, Denoising, Signal Retrieval.

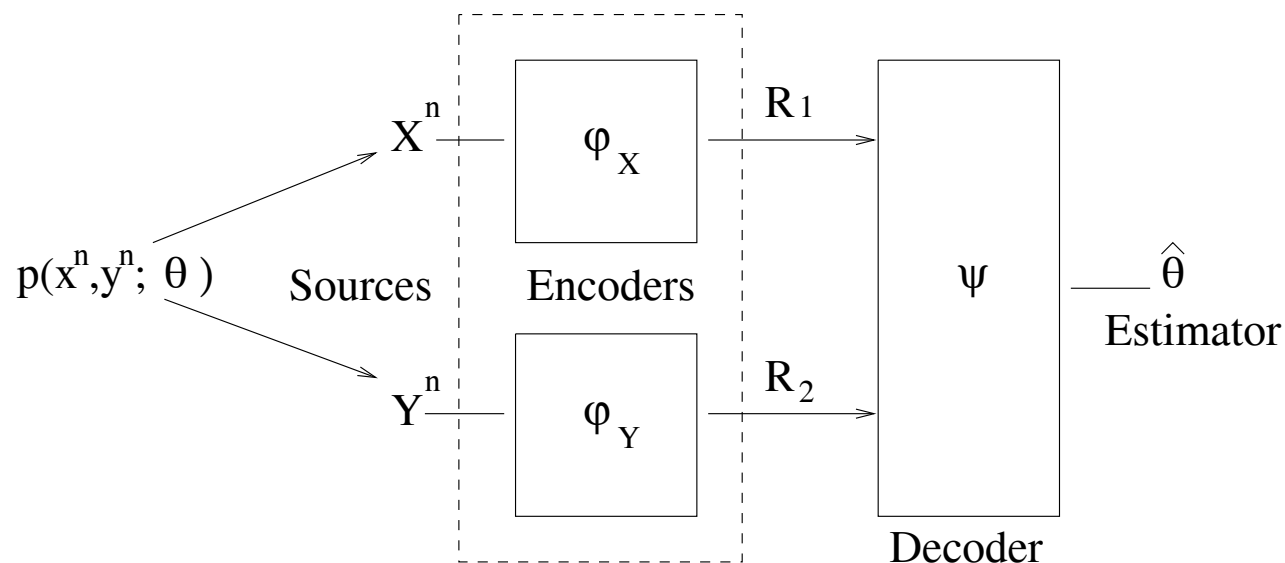
Multiterminal Source Coding

Multiterminal Estimation is by comparison much less studied. Two approaches exist:

-Zhang and Berger (1988)

-Han and Amari (1995,1998).

Problem Setup



The data: $(X^n, Y^n) \sim p_\theta(x^n, y^n)$, where $\mathcal{X} = 1, \dots, \mathcal{K}$ and $\mathcal{Y} = 1, \dots, \mathcal{L}$

Encoders: $\varphi_X : X^n \rightarrow \varphi_X(X^n)$, where $\frac{1}{n} \log_2 \|\varphi_X\| \leq R_1 + \delta$

$\varphi_Y : Y^n \rightarrow \varphi_Y(Y^n)$, where $\frac{1}{n} \log_2 \|\varphi_Y\| \leq R_2 + \delta$

Decoder: $\psi(\varphi_X(X^n), \varphi_Y(Y^n))$ Estimator: $\hat{\theta}(\psi(\varphi_X(X^n), \varphi_Y(Y^n)))$

Problem: $\min_{\varphi_X, \varphi_Y, \psi, \hat{\theta}} \lim_{n \rightarrow \infty} nV(\hat{\theta}(\psi(\varphi_X(X^n), \varphi_Y(Y^n)))) = V(\theta | R_1, R_2)$

Lossless/Lossy Compression and Estimation

Data: Observe sample (x^n, y^n) with marginal and joint types denoted by $\lambda_x, \lambda_y, \lambda_{xy}$.

Lossless: estimator $\tilde{\theta}(x^n, y^n) = \tilde{\theta}(\lambda_x, \lambda_y, \lambda_{xy})$, with types restricted to $k = 1, \dots, K$ and $l = 1, \dots, L$ (m-coordinates).

Lossy: Defined by a *stochastic mapping*,
test-channels $\stackrel{\text{def}}{=} \begin{aligned} p_\theta(u|x) &: \mathcal{X} \rightarrow \mathcal{U} \\ p_\theta(v|y) &: \mathcal{Y} \rightarrow \mathcal{V} \end{aligned}$

where $\mathcal{U} = 0, \dots, N$, $\mathcal{V} = 0, \dots, M$. $U \rightarrow X \rightarrow Y \rightarrow V$ form a Markov chain.

Test-channel output: the *codewords* U^n, V^n

Compression ctd

- Separate compression: $p_{\theta}(u, v|x, y) = p_{\theta}(u|x)p_{\theta}(v|y)$
- θ is unknown, can't have the test-channels depend on it
- fixed test-channels inefficient approach

Universal Coding scheme, fix the test-channels for each data type:

$$p_{\lambda_x}(u|x), p_{\lambda_y}(v|y)$$

For large n , we can approximate the test-channels by the *conditional types*: $\tilde{p}_{\lambda_x}(u|x) = \frac{\lambda_{ux}}{\lambda_x}$, $\tilde{p}_{\lambda_y}(v|y) = \frac{\lambda_{vy}}{\lambda_y}$

Lossy Compression:

$$(X^n, Y^n) \xrightarrow{R_1, R_2} (U^n, V^n) \stackrel{m\text{-coord}}{=} (\lambda_u, \lambda_v, \lambda_{uv}) \\ \xrightarrow{0\text{-rate}} (\lambda_x, \lambda_y)$$

Estimation based on lossy compression: $\hat{\theta}(\lambda_x, \lambda_y, \lambda_{uv})$

The Problem: Find bounds on asymptotic efficiency of
 $\hat{\theta}(\lambda_x, \lambda_y, \lambda_{uv})$

Zhang and Berger (1988)

- Encoder ensemble moments for additive estimators
- Restrict $\hat{\theta}(\lambda_x, \lambda_y, \lambda_{uv})$ to a linear functional of $\tilde{\theta}(\lambda_x, \lambda_y, \lambda_{xy})$
- Simplicity: single-letter characterization

Han and Amari (1995, 1998)

- Asymptotic distribution of $(\lambda_x, \lambda_y, \lambda_{uv}) \rightarrow$ MLE, Cramer-Rao bound
- Universal *decoding* argument
- Complexity: uncomputable for sources $>$ binary

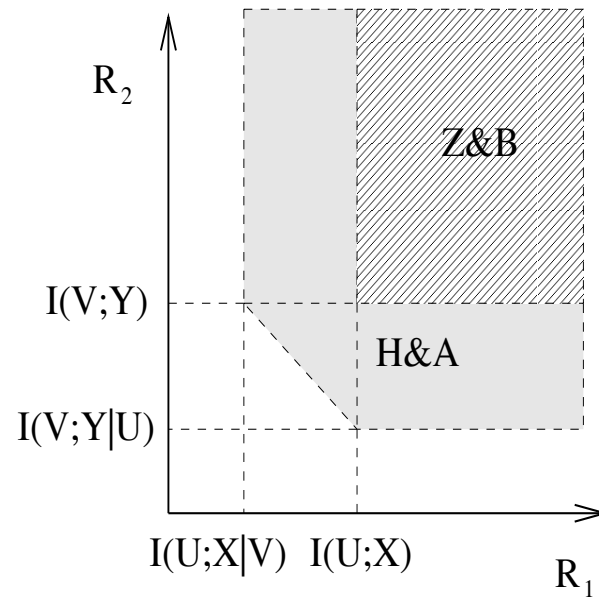
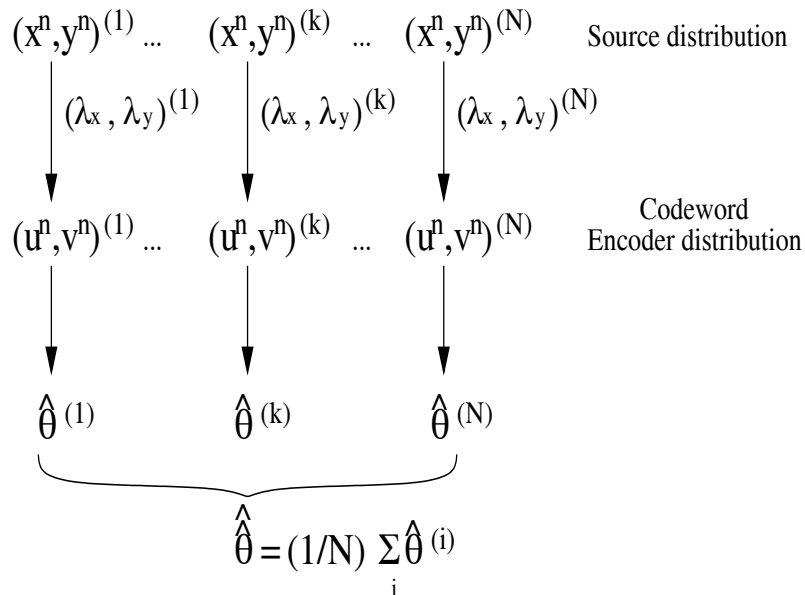
Zhang and Berger

There exists a code (f, g) with rates (R_1, R_2) s.t.

$$nV[\hat{\theta}_{\lambda_x, \lambda_y}(f(X^n), g(Y^n))] \leq V[\hat{\theta}(U, V)] + E[E(\tilde{\theta}(X, Y)|X)]^2 + E[E(\tilde{\theta}(X, Y)|Y)]^2 - E[E(\hat{\theta}(U, V)|UX)]^2 - E[E(\hat{\theta}(U, V)|VY)]^2 + O(n^{-1/2})$$

Rate compatibility conditions: $R_1 \geq I(U; X)$ and $R_2 \geq I(V; Y)$

Key: Compute *ensemble moments* over random codebooks.



Han and Amari 1995,1998

$(\lambda_x, \lambda_y, \lambda_{uv}) = \bar{\Lambda}_f$: HA show $\bar{\Lambda}_f \underset{n \rightarrow \infty}{\sim} N(\bar{p}_{\theta,f}, H_f G H_f^T)$

There exists a code (f, g) with rates R_1, R_2 and a decoder h s.t.

$$nV[\hat{\theta}_{\lambda_x, \lambda_y}(h(f(X^n), g(Y^n)))] \leq \left\{ \nabla \bar{p}_{\theta,f} (H_f G H_f^T)^{-1} \nabla \bar{p}_{\theta,f}^T \right\}^{-1} + O(n^{-1/2})$$

Rate compatibility conditions:

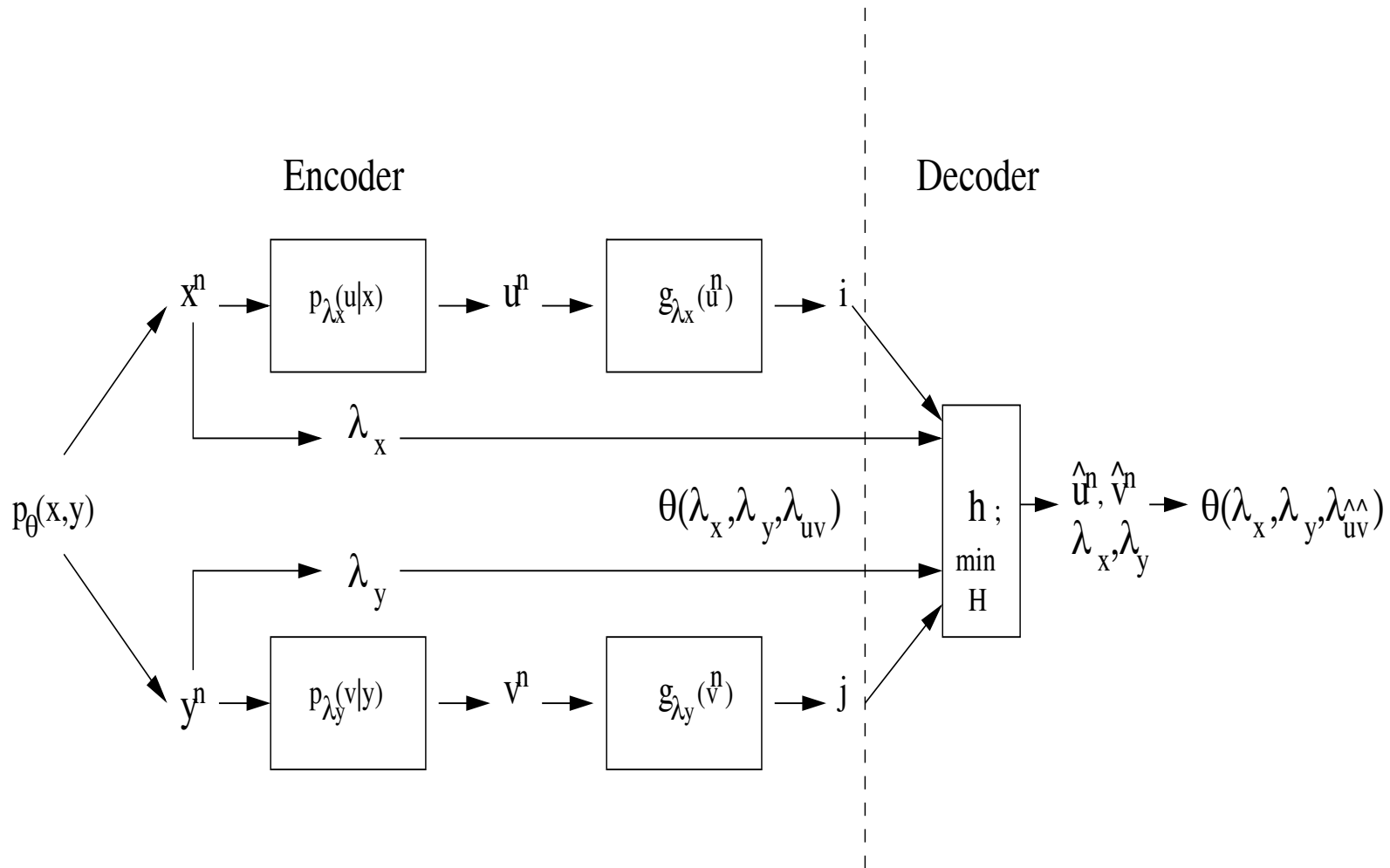
$$R_1 \geq I(U; X|Y), R_2 \geq I(V; Y|X), R_1 + R_2 \geq I(UV; XY)$$

Codebook constraint on type distribution Λ_f enters via projection

operator $H_f : \sum_{u x y v} \lambda_{u x y v} = 1, \lambda_{u x} = \tilde{p}(u|x)\lambda_x, \lambda_{v y} = \tilde{p}(v|y)\lambda_y$

Computing H_f in terms of the m-coordinates is a non-trivial exercise.

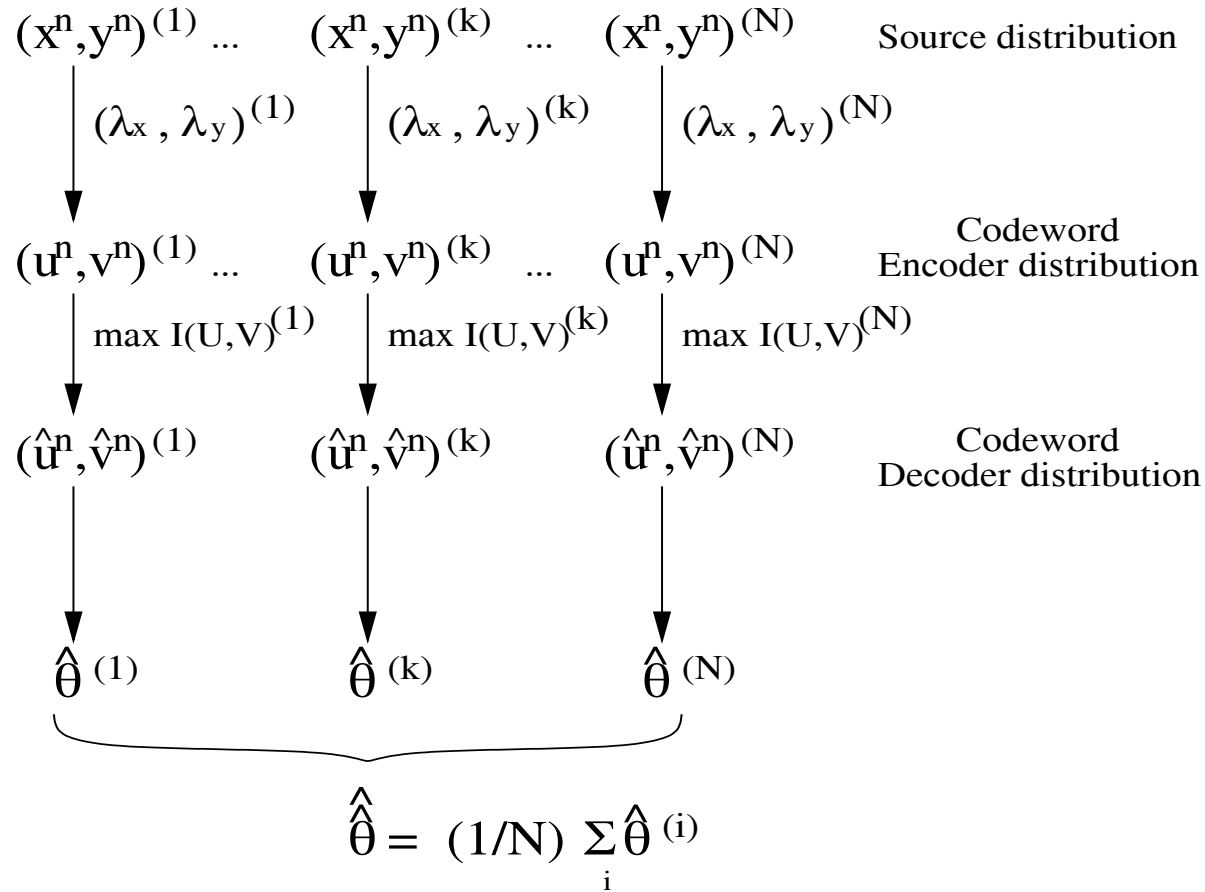
Han and Amari



Extending Zhang and Berger

- Push *Ensemble moment* computation to the decoder side.
- Use test-channels that map from/to exponential families.
- Consider estimates $\hat{\theta}$ that are “nice” functionals of mean-value parameters (additive and asympt. unbiased MLEs)
- A delta-method argument gives an *approximation* of the Cramer-Rao bound of Han and Amari, and
 1. it's *computable* for large source alphabets
 2. it has a *single-letter characterization* (suffices to compute moments under $p_{\theta}(u, x, y, v)$, cmp. $p_{\theta}(\Lambda_f)$).

Extending Zhang and Berger



New upper bound on $V(\theta|R_1, R_2)$

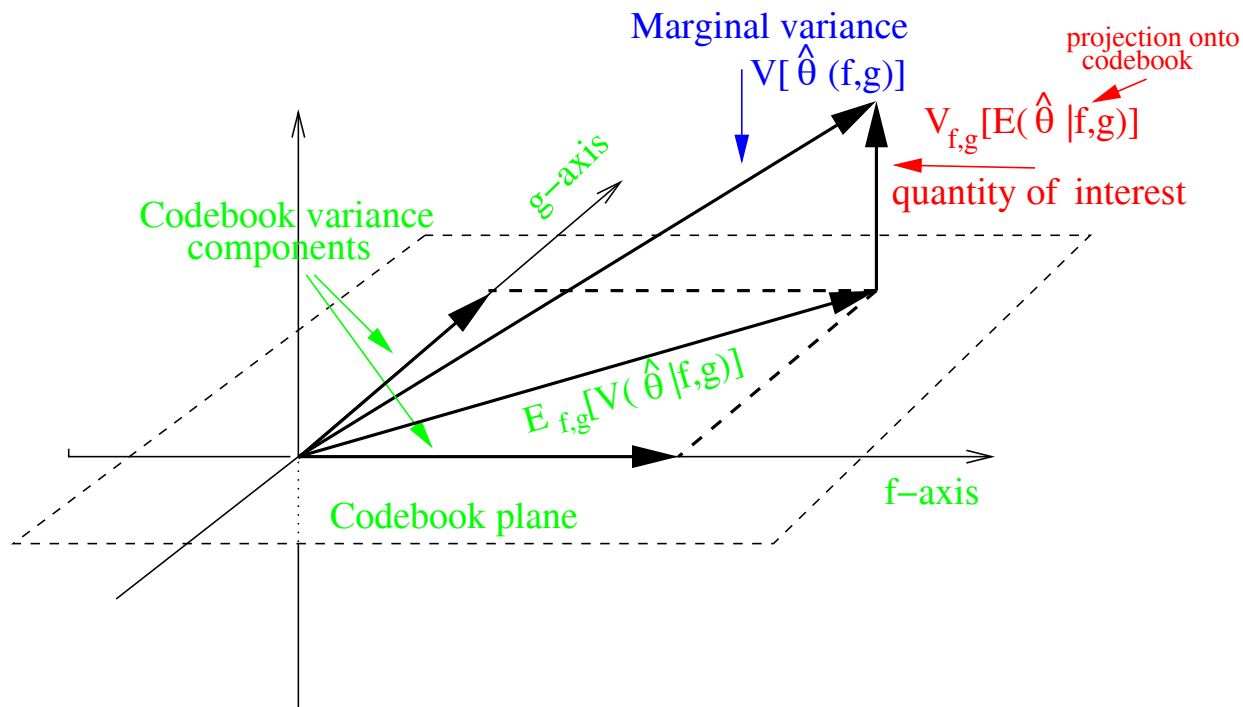
There exists a code (f, g) with rates R_1 and R_2 and a decoder h s.t.

$$\begin{aligned}
 nV[\hat{\theta}(h(f(X^n), g(Y^n)))] &\leq V[\hat{\theta}(U, V)] \\
 &\quad - E[E(\hat{\theta}(U, V)|UX) - E(\hat{\theta}(U, V)|X)]^2 \\
 &\quad - E[E(\hat{\theta}(U, V)|VY) - E(\hat{\theta}(U, V)|Y)]^2 + O(n^{-1/2}) \\
 \underset{\substack{\text{orthog. codebooks} \\ \text{exp. family}}}{=} & V[\hat{\theta}(U, V)] - E_{XY}(V_{UV|XY}[\hat{\theta}(U, V)]) \\
 \underset{\substack{\text{variance} \\ \text{decomp.}}}{=} & V_{XY}(E_{UV|XY}[\hat{\theta}(U, V)])
 \end{aligned}$$

Rate compatibility conditions: same as Han and Amari.

Geometric interpretation

$$nV[\hat{\theta}(h(f(X^n), g(Y^n))) \leq V[\hat{\theta}(U, V)] - E_{XY}(V_{UV|XY}[\hat{\theta}(U, V)])$$



Or in codebook notation,

$$V[\hat{\theta}(f, g)] - E_{f, g}[V(\hat{\theta}(f, g)|f, g)] = V_{f, g}[E(\hat{\theta}(f, g)|f, g)]$$

Estimation of the crosscorrelation between two separately compressed sources.

Bivariate gaussian model: $p_{\theta}(x^n, y^n), \theta = (\sigma_x^2, \sigma_y^2, \rho)$

Test-channels: $p(u|x) \sim N(x, \sigma_n^2)$

$p(v|y) \sim N(x, \sigma_m^2)$

Bound still applicable for cts sources if we *discretize* (U, X, Y, V) .

New bound $\implies nV(\hat{\rho}(h(f(X^n)), g(Y^n))) \leq$

$$(1 - \rho^2)^2 + \left(\frac{1}{2^{2R_1} - 1} + \frac{1}{2^{2R_2} - 1}\right) + \left(\frac{1}{2^{2R_1} - 1} \frac{1}{2^{2R_2} - 1}\right) \\ - \rho^2 \left(\frac{1}{2^{2R_1} - 1} + \frac{1}{2^{2R_2} - 1} + \frac{1}{2^{2R_2}(2^{2R_1} - 1)} + \frac{1}{2^{2R_1}(2^{2R_2} - 1)}\right)$$

Compare with Zhang and Berger (stronger rate constraints):

$$nV(\hat{\rho}(f(X^n), g(Y^n))) \leq (1 + \rho^2) + (1 - \rho^2) \left[\frac{1}{2^{2R_1} - 1} + \frac{1}{2^{2R_2} - 1}\right] \\ + \frac{1}{2^{2R_1} - 1} \frac{1}{2^{2R_2} - 1}$$

Conclusion

- Simple extension of Zhang and Bergers approach
- Same rate compatibility conditions as Han and Amari
- Computable for most sources
- Has a geometric interpretation

Future Work:

- “Nice” functionals may not correspond to the optimal estimators $\hat{\theta}$
- Best approximating exponential family
- Dependent case